Does Local Government Cooperation promote efficiency gains? Evidence from Italian Municipal Unions

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- In the last 50 years municipalities across Europe have been affected by economic and government budget challenges
- \longrightarrow pressure on performances in terms of efficiency, effectiveness and quality of public services.

Two issues:

- 1. general increase in the demand of public goods
 - Citizens more conscious and demand more skilled set of public goods, and higher accountability than in the past.
- 2. stricter public finance requirements imposed by EU to local governments
 - \longrightarrow municipalities had to reduce their expenditures.

Consequence:

- difficulties for small municipalities to meet the demands of local public goods while reducing their expenditure
- territorial scale no longer compatible with the increasing scale of production required to provide efficiently high standard level of public services

Possible solutions: to rethink service delivery of services through new different institutional tools

- amalgamation of municipalities
- inter-municipal cooperation.

- Compulsory merging of neighbouring borders and creation of new entities.
- Aim to reduce the number of units of sub-national governments,
- and to achieve efficiency gains, from both the exploitation of economies of scale
 - larger area can be served after the agreements,
- internalisation of externalities (Oates, 1972; Case et al., 1993).

However, the municipal amalgamation is often difficult to achieve, especially due to political resistance of the local policy maker to merge (Mello and Lago-Penas, 2013).

Inter-municipal cooperation

- Governance structure where municipalities reciprocally cooperate
- to provide a wide range of public services or organize service delivery between partners.
- Municipalities can transfer their own decision-making powers on some public services to a standing organization.
 - \longrightarrow new entity which provides only the transferred public services.
- Thus municipalities can co-operate without being replaced by the standing organization.
- Inter-municipal cooperation is a more flexible solution than amalgamation (Dollery et al., 2006; Feiock and Scholz, 2009; Mello and Lago-Penas, 2013; Blaeschke, 2014)
- Municipalities can maintain, on their own territory, local political representatives
- and decide on their own fiscal policy.

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- Saarimaa and Tukiainen (2015), using Finnish data, find that municipalities, before amalgamation, shift part of the costs of additional expenditures to the future partners
 - \longrightarrow increasing debt or liquidating assets.
- Fritz and Feld (2015) also observe higher debt dynamics for a sample of amalgamated German municipalities, in particular
 - if the number of amalgamated municipalities increases,
 - and if municipalities are either forced to merge or annexed.

- Many empirical works on the determinants of the inter-municipal cooperation.
- In particular, they find that relevant factors in favour of inter-municipal cooperation are
 - size of municipalities (Brasington, 2003; Carr et al., 2007)
 - regional characteristics (Feiock, 2007; LeRoux and Carr, 2007)
 - geographic factors (Morgan and Hirlinger, 1991; Post, 2002)
 - fiscal revenue (Di Porto et al., 2013)
 - spatial proximity of municipalities (Di Porto et al., 2016)

- Few studies on the impact of inter-municipal cooperation on socio-economic variables and
- no studies that empirically explore the ex-post impact of the inter-municipal cooperation on both financial and service outcomes.
- The aim of this paper is to fill this gap in the literature, by studying the causal impact of inter-municipal cooperation on local spending and on the provision of the local public services.

- We investigate the Italian experience in the inter-municipal cooperation process which starts in 1990 with the institution of the Municipal Union (Unione dei Comuni).
- We use unique administrative data on municipalities in the region Emilia Romagna over the period 2001-2011.
- By exploiting the different timing in entering/forming Municipal Unions, we employ difference-in-differences models to identify the causal effect of the inter-municipal cooperation.
- We adopt parametric and non-parametric difference-in-differences matching models to control for several sources of sample bias.

- Being a member of a Municipal Union reduces total current per-capita expenditures by around 5%, compared to municipalities not in a Union.
- We find that the expenditure reduction is consistent and increasing up to nine years after joining a Municipal Union.
 - \implies persistency of the policy effect.
- Results robust to a large number of checks
- We also find that spending cuts are not associated with a downsizing of local services, which confirms that the Union is effectively increasing municipalities efficiency.

- In Italy, 8,000 municipalities and, approximately 70% of them have a population lower than 5,000 inhabitants.
- The Law 142/1990 has formally introduced the inter-municipal cooperation.
- Municipalities can transfer their own decision-making powers, in terms of expenditure decisions, to a standing organization called Unione di Comuni.
- Municipalities transfer the money related to the public function(s) they want to share, and the Union provides the corresponding service(s).

- Union is a legal entity,
 - with its own balance sheet,
 - its own president chosen among the mayors of municipalities joining the Union
 - its own council composed by the council members of cooperating municipalities.
- Each municipality can be member of only one Union.
- The functions commonly transferred are: Administration and Management, Municipal police, Education, Roads & Transport Services, Planning and Environment and Social welfare.

- Municipalities can also transfer Economic development, In-house production services, Culture, Sport and Tourism.
- Each region, through its own law, can stimulate and promote Municipal Unions within its territory, by means of regional transfers.
- Some region used financial incentives (Veneto, Toscana and Emilia Romagna).
- Lombardia has created a special register of Municipal Unions (Unioni di Comuni Lombardi), such that only registered municipalities have access to regional transfers.

- In 2007 total expenditures of Municipal Unions about 0.10% (403 millions of euro) of total local expenditures in Italy (350 billions of euro).
- In 2013, the total expenditures of Municipal Unions more than doubled, 0.30% (970 million of euro) of the total local expenditures in Italy (334 billion of euro).
- Revenues of Municipal Unions: transfers from municipalities within the Union and transfers from State and regional governments.

Empirical Approach

- The organisation process of Municipal Unions has not been homogeneous across regions in Italy, and over time (during the period of our analysis 2001-2011).
- Municipalities in Unions located in different regions are not properly comparable
 - \implies we cannot identify a unique (aggregate) effect of the policy on local expenditures.
- Analysis restricted to Emilia Romagna: one of the richest and more populated Italian region
 - average population of about 4 million inhabitants over the period 2001-2011 (approximately 7.50% of the Italian population)
 - average GDP, over the same period, is 116 billion euros (approximately 9% of the Italian GDP).
 - inter-municipal cooperation is a widespread phenomenon: number of Municipal Union has noticeably increased.

Data

Combination of different archives from

- Italian Ministry of the Interior,
- the Italian Ministry of Economy
- and the Italian Institute of Statistic.

Information on:

- municipal financial data
- municipal demographic and socio-economic data: total current expenditures, population size, age structure, average income of inhabitants.

In Emilia Romagna

- 348 municipalities for the period 2001-2011
- we exclude Bologna, a Metropolitan Area with a wider range of services

Final sample:

- 335 municipalities, over the period 2001-2011
- a total of 3,686 observations.

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Data definition and sources

| Variable | Definition and measure | Available from-to | Source |
|--------------------------------------|--|-------------------|--|
| Log expenditure | Log of current expenditure per resident; 2011 Euros | 2001-2011 | Italian Ministry of Interior |
| Municipal Union | Dummy variable that takes on the value one if municipality i at time t belongs to a Municipal Union and zero otherwise. | 2001-2011 | Italian Ministry of Interior |
| Birth rate per capita | birth rate per capita | 2001-2011 | ISTAT |
| Number of childre enrolled in infant | Number of childre enrolled in infant school (per-capita) | 2004-2011 | ISTAT |
| Net migration | Difference between new regsitred members and unregistered members | 2002-2010 | ISTAT |
| Per capita road car crash | Number of accients within the muncipal roads | 2001-2011 | ISTAT |
| Taxpayers | share of the taxpayesr of the municipality | 2001-2011 | Italian Ministry of Economy, Department of Finance |
| population | Population of the municipality | 2001-2011 | ISTAT |
| child | Share of the population aged between 0-5 | 2001-2011 | ISTAT |
| bld | Share of the population over the age of 65 | 2001-2011 | ISTAT |
| population density | Numbers of citizens per area | 2001-2011 | Our computation |
| income | Real personal income tax base per resident; 2011 Euros | 2001-2011 | Italian Ministry of Economy, Department of Finance |
| l/population | inverse of the population | 2001-2011 | Our computation |
| Revenue of Muncipal Unions | Sum of the revenue from fees and charges and transfers from other level of governmenrs per resident; 2011 Euros | 2001-2011 | Italian Ministry of Interior |
| Permanence | Number of years joining the Municpal Union | 2001-2011 | Our computation |
| Permanence square | Square of number of years joining the Municpal Union | 2001-2011 | Our computation |

Descriptive statistics

Table: Municipalities in Unions in Emilia Romagna 2001-2011

| Variables | Obs | Mean | Std. Dev. | Min | Max |
|---------------------------------|------|-----------|-----------|-----------|------------|
| Log expenditure | 3686 | 6.663 | 0.287 | 5.733 | 8.329 |
| Municipal Union | 3784 | 0.196 | 0.397 | 0.000 | 1.000 |
| Population | 3686 | 11419.580 | 23150.210 | 91.000 | 186690.000 |
| Child | 3686 | 0.051 | 0.013 | 0.000 | 0.083 |
| Aged | 3686 | 0.238 | 0.066 | 0.120 | 0.638 |
| Income | 3686 | 13223.430 | 2086.341 | 5425.244 | 20525.250 |
| Population density | 3686 | 0.017 | 0.025 | 0.000 | 0.275 |
| Taxpayers | 3686 | 0.706 | 0.099 | 0.324 | 1.083 |
| 1/population | 3686 | 0.000 | 0.001 | 0.000 | 0.011 |
| Revenue of Municipal Unions | 3684 | 11.198 | 37.796 | 0.000 | 321.683 |
| Permanence | 3784 | 0.587 | 1.812 | 0.000 | 15.000 |
| Permanence square | 3784 | 3.627 | 16.803 | 0.000 | 225.000 |
| Birth rate per capita | 3686 | 0.009 | 0.003 | 0.000 | 0.027 |
| N. children in infant school pc | 2351 | 0.024 | 0.007 | 0.000 | 0.061 |
| Net migration | 3784 | 41.461 | 121.474 | -1773.000 | 1366.000 |
| Per capita road car crash | 3592 | 0.004 | 0.003 | 0.000 | 0.024 |

Figure: Emilia Romagna municipalities - 2001



• Only one MU, including 9 municipalities and serving 20,767 inhabitants, around 1% of the regional population.

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Union

Figure: Emilia Romagna municipalities - 2011



• 31 MU, over 160 municipalities and 1.5 million of inhabitants, that is 34% of the total population of Emilia Romagna.

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- On average 5 municipalities per Union (from a minimum of 2 to a maximum of 10)
- and average population of Union is 43,000 inhabitants.
- Emilia Romagna approved the LR 2008, n.10 aiming at rationalising public expenditures through a reorganisation of its territory.
 - transformed mountain communities in MU
 - encouraged municipalities to form Unions, giving direct financial incentives.

Figure: Percentage of municipalities in union



- The rate of participation of municipalities in Municipal Unions is less than 20% until 2007
- striking increase from 2008, reaching almost 50% by the end of the decade
- by 2016 the percentage of municipalities in unions is 81%

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| Table: Municipalities in | Unions in | Emilia | Romagna | 2001-2011 |
|--------------------------|-----------|--------|---------|-----------|
|--------------------------|-----------|--------|---------|-----------|

| Year | Municipal Unions | Municipalities not in Unions | Municipalities in Unions | % of Municipalities in Unions | Total Municipalities |
|------|------------------|------------------------------|--------------------------|-------------------------------|----------------------|
| 2001 | 1 | 328 | 9 | 2.67 | 337 |
| 2002 | 5 | 314 | 23 | 6.82 | 337 |
| 2003 | 5 | 303 | 23 | 7.06 | 326 |
| 2004 | 6 | 305 | 32 | 9.50 | 337 |
| 2005 | 6 | 303 | 32 | 9.55 | 335 |
| 2006 | 10 | 277 | 52 | 15.81 | 329 |
| 2007 | 10 | 281 | 54 | 16.12 | 335 |
| 2008 | 13 | 266 | 70 | 20.83 | 336 |
| 2009 | 25 | 199 | 132 | 39.88 | 331 |
| 2010 | 29 | 193 | 150 | 43.73 | 343 |
| 2011 | 31 | 180 | 160 | 47.06 | 340 |

- Dependent variable: total current expenditure of each municipality, in per-capita terms.
- It is used as an aggregate measure of performances of municipalities in Union and not.
- For municipalities in Union the measure includes their transfers to the Union
 - indeed the services provided by the Unions are financed by transfers and by direct expenditures.
- The average pc expenditure of municipalities in U is 751.10 euros whereas for the others is 835.99 euros. Differences of 84.88 euros



Objective: identify the causal effect of being a member of a MU on the spending decisions of single M.

- Ideal framework: to compare decisions on expenditure for M in a U (treated group), to the same decisions for M in the counterfactual situation of not being in the U.
- This is impossible.

- Best alternative is a randomized control trial: to assign participation and non participation in a MU across municipalities, and compare the average expenditures of the two groups.
- Our analysis: quasi-experimental approach
 - define a suitable control group that can credibly estimate the counterfactual.
 - Main issue: unobservable characteristics between M in U and NU, and which might be correlated to the expenditures.
 - Methodology used difference-in-differences \longrightarrow remove unobservables fixed over time.

- For each year, we have municipalities in Unions (treated group, TG) and municipalities not in Unions (control group, CG).
- We compare the change in expenditures in TG before and after the participation in a MU, to the change in expenditures in CG for the same period.
- We estimate the following two-way fixed effect linear model

$$Y_{it} = \mu_i + \tau_t + \gamma M U_{it} + \beta \mathbf{x_{it}} + \varepsilon_{it}$$
(1)

- Y_{it} is log per capita expenditure in municipality *i* at time *t*
- *MU_{it}*: dummy variable =1 if municipality *i* at time *t* belongs to the Municipal Union and zero otherwise.
- μ_i municipalities fixed effects
- τ_t exogenous shocks, common to all municipalities in period t.
- $\boldsymbol{x_{it}}$: time-varying demographic and socio-economic characteristics.
- γ is the difference-in-difference estimate of the effect of being in a Municipal Union on expenditure.
- standard errors clustered at municipal level.

Control variables

- Population of M, population density (number of citizens per area), and inverse of population.
 - variables that capture scale economies or dis-economies in the provision of public goods and congestion effects.
- Proportion of citizens aged 0-5, proportion of citizens aged over 65
 - account for some specific public needs (e.g., nursery school, nursing homes for the elderly).
- Average per capita income proxied by the personal income tax base, and the proportion of taxpayers.
- Total per capita revenue of MU: own total revenue + transfers from higher levels of government.
 - variable that allows to separate the effect of being in MU from the variation in the financial resources raised by the Municipal Union.

- We evaluate whether the effect of the policy varies with respect to the permanence in MU.
- We take into account the length of time spent by each municipality in the Union after joining it.
- This is an important issue since we are dealing with multiple treatment groups and multiple time periods.
- We estimate a modified version of equation 1, where we add a continuous variable that measures the permanence in the Union (from zero to 11 years), and we also include its quadratic term.

Sample selection bias

Two potential sources of bias

- 1. effect of entering a Union is not homogeneous
 - This bias arises when municipalities in the TG are somehow different than those in CG.
- 2. the effect varies according to the characteristics of the municipalities.
 - this bias might be due to different distributions, within the treatment and the control groups, in the vector of observable characteristics that affect expenditures.
- Solution: propensity score matching models.
 - The main purpose is to find a group of non-treated municipalities, who are similar to the treated in all relevant pre-treatment characteristics
 - the only remaining difference being that one group enters a Union and another group does not.

1. Estimation of the propensity score

- The probability of entering a U conditional on pre-treatment characteristics x, P(x) = Pr(MU = 1|x)
- We use data from the 2001 Census and run a logit regression,
- \longrightarrow dependent variable: dummy =1 if a municipality in MU between 2001-2011 and zero otherwise.
 - control variables: municipality located close to the coast, rural municipality, surface in square km of the municipality, municipality located in plain, on hills, and in mountains; municipal unemployment rate; number of houses; number of firms; interaction between surface and houses; interaction number of firms and unemployment rate.

- 2. Once we have obtained the (PS), following Smith and Todd (2005), we adopt a trimming procedure to define the common support
 - region of values of PS that have positive density within both the treatment and control groups distributions.
- 3. We re-estimate equation 1 by using information only on the observations that lie on the common support.

DiD propensity score matching I

- Within the common support, the distribution of **x** might be different between treated and control observations, keeping the second source of bias.
- We control for it by using a non-parametric DiD kernel matching approach.
- We choose 2008 and 2010 as pre-treatment and pos-treatment period, respectively.
- Following Heckman et al. (1998) we estimate

$$\gamma^{DiD} = \sum_{i \in MU} \left\{ \left[Y_{it_1} - Y_{it_0} \right] - \sum_{j \in NMU} W_{ij} \left[Y_{jt_1} - Y_{jt_0} \right] \right\} w_i \quad (2)$$

• *W_{ij}* and *w_i* are weights to construct the counterfactual and re-weighting the treated sample, respectively.

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|--------------------------|-------|---------|-------------|---------|

- w_i is the re-weighting for the treated sample.
- t_0 and t_1 are time periods before and after entering a Union.
- MU: municipalities not in U in t_0 that will join a Union in t_1
 - we include only municipalities that join a U in 2009.
- *NMU*: municipalities not in U in t_0 that will remain out in t_1
 - municipalities that never joined a Union.
- Y is the expenditure of municipalities.
 - We perform the matching approach by comparing expenditure between municipalities in TG and CG in 2008 and in 2010.
| Dependent variable: | Log Expenditures | | | | | | | | |
|------------------------|------------------|--------------|--------------|----------------|--------------|--------------|------------------------------|--|--|
| | | Full sample | | Matched sample | | | Kernel matching ^a | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | | |
| Municipal Union | -0.065*** | -0.047*** | -0.030** | -0.063*** | -0.048*** | -0.030** | | | |
| | (0.012) | (0.015) | (0.015) | (0.013) | (0.016) | (0.015) | | | |
| Permanence | | | -0.022*** | | | -0.021*** | | | |
| | | | (0.006) | | | (0.006) | | | |
| Permanence square | | | 0.002*** | | | 0.002*** | | | |
| | | | (0.000) | | | (0.000) | | | |
| Union ₀₈₋₁₀ | | | | | | | -0.069*** | | |
| | | | | | | | (0.020) | | |
| N | 3686 | 3586 | 3586 | 3411 | 3311 | 3311 | 434 | | |
| Year FE | \checkmark | \checkmark | \checkmark | \checkmark | \checkmark | \checkmark | \checkmark | | |
| Municipality FE | \checkmark | \checkmark | \checkmark | \checkmark | \checkmark | \checkmark | | | |
| Municipality controls | | \checkmark | \checkmark | | \checkmark | \checkmark | | | |

Table: Effect of the Union on Log Expenditures

Standard errors clustered at municipality level.

Significance levels: *** 1% ** 5% * 10%.

^a Kernel difference in differences matching. Control group= municipalities never in Union.

Treatment group: municipalities that joined a Union in 2009.

Sample size restricted to years before and after the Union 2008 and 2010

Table: Permanence in the municipal union

| Dependent variable: | Log Expendi | tures - 2001-2011 |
|------------------------|-------------|-------------------|
| | Full sample | Matched sample |
| | (1) | (2) |
| 1 year | -0.030** | -0.030** |
| | (0.015) | (0.015) |
| 2 years | -0.049*** | -0.049*** |
| | (0.016) | (0.016) |
| 3 years | -0.064*** | -0.064*** |
| | (0.018) | (0.018) |
| 4 years | -0.075*** | -0.074*** |
| | (0.020) | (0.021) |
| 5 years | -0.081*** | -0.079*** |
| | (0.023) | (0.024) |
| 6 years | -0.082*** | -0.080*** |
| | (0.026) | (0.027) |
| 7 years | -0.080*** | -0.077*** |
| | (0.029) | (0.029) |
| 8 years | -0.072** | -0.069** |
| | (0.032) | (0.032) |
| 9 years | -0.061* | -0.057 |
| | (0.034) | (0.035) |
| 10 years | -0.044 | -0.040 |
| | (0.037) | (0.038) |
| 11 years | -0.024 | -0.018 |
| | (0.040) | (0.041) |
| Significance levels: * | ** 1% ** 5% | * 10%. |
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Analysis on different outcomes I

- The results show that inter-municipal cooperation increases the efficiency of the single municipality,
 - i.e. its expenditure reduces once the municipality enters a Union.
- However, these savings might not be a real gain in efficiency
- but the consequence of a reduction of the level or the quality of public services.

We test whether local public services are affected by inter-municipal cooperation using four proxies of their level.

- 1. per capita birth rate,
- 2. net migration to the municipality
- 3. p.c. primary school class size
- 4. p.c. number of road accidents.

If local services were to decline we would expect a negative impact on these indicators, as a consequence of the reduction in expenditures. For example:

- poor local public services may imply lower migration and lower birth rates, as a result of lower attractiveness of the municipality.
- Less expenditures may increase the school class size, a typical indicator of the quality of the school service.
- A decrease in current expenditures which leads to a worsening of roads or bridges maintenance might increase car accidents.

Table: Effect of the Union on alternative output

| | (1) | (2) | (3) | (4) |
|-------------------------|--------------|--------------|--------------|--------------|
| Full sample | () | () | () | () |
| Municipal Union | 0.000 | 0.000 | -6.256 | -0.000 |
| | (0.000) | (0.000) | (7.666) | (0.000) |
| Ν | 3613 | 2320 | 3613 | 3524 |
| Matched sample | | | | |
| Municipal Union | 0.000 | 0.000 | -4.414 | -0.000* |
| | (0.000) | (0.000) | (6.984) | (0.000) |
| N | 3586 | 2301 | 3586 | 3498 |
| Year FE | \checkmark | \checkmark | \checkmark | \checkmark |
| Municipality FE | \checkmark | \checkmark | \checkmark | \checkmark |
| Municipality controls | \checkmark | \checkmark | \checkmark | \checkmark |
| Standard errors cluste | red at mu | nicipality l | evel. | |
| Significance levels: ** | * 1% ** 5 | 5% * 10%. | | |
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- It is clear that being in U does not affect any of our four local service indicators.
- Inter-municipal cooperation is not associated to any reduction in the provision of public services
- Expenditure savings coming from the participation to a U can be interpreted as an efficiency gain.

- To verify the DiD key identification assumption of common time trend in the pre-treatment periods for both comparison groups.
- We perform a test re-estimating equation 1 including the interactions of the time dummies and the treatment indicator for the first three pre-treatment periods → "leads"
- If the expenditure trends between TG and CG are the same, then the coefficients of the interactions should be insignificant,
 - the difference in differences is not significantly different between the two groups in the pre-treatment period.

- Additionally, we include the interaction of the time dummies after the treatment (up to 3 years) with the treatment indicator → "lags"
 - this can show whether the treatment effect changes over time.

We estimate the following version of equation 1:

$$Y_{it} = \mu_i + \tau_t + \sum_{j=0}^m \gamma_{-j} M U_{i,t-j} + \sum_{j=1}^n \gamma_{+j} M U_{i,t+j} + \beta \mathbf{x_{it}} + \varepsilon_{it}$$
(3)

- the sum of γ_{+j} allows for n leads or anticipatory effects.
- A test of the difference in differences assumption is $\gamma_{+j}=0$ for each $j=1\ldots n$

- the coefficients of all leads of the treatment should be zero.
- the sum of γ_{-j} allows for m lags effects
- γ_{-j} may not be identical and if the effect of the treatment is growing over time γ_{-j} increases in j.

| Dependent variable: | Log Expenditures | | | | | | |
|-----------------------|------------------|--------------|----------------|--------------|--|--|--|
| | | | | | | | |
| | Full sample | | Matched sample | | | | |
| | (1) (2) | | (3) | (4) | | | |
| Municipal Union | -0.044*** | -0.018* | -0.043*** | -0.017* | | | |
| | (0.011) | (0.009) | (0.011) | (0.009) | | | |
| $lead_{t+1}$ | 0.006 | 0.004 | 0.007 | 0.005 | | | |
| | (0.010) | (0.010) | (0.010) | (0.010) | | | |
| $lead_{t+2}$ | -0.010 | -0.009 | -0.013 | -0.013 | | | |
| | (0.011) | (0.011) | (0.011) | (0.011) | | | |
| $lead_{t+3}$ | 0.013 | 0.008 | 0.015 | 0.010 | | | |
| | (0.009) | (0.010) | (0.009) | (0.010) | | | |
| lag_{t-1} | | -0.025** | | -0.025** | | | |
| | | (0.010) | | (0.010) | | | |
| lag_{t-2} | | -0.022* | | -0.021 | | | |
| | | (0.013) | | (0.013) | | | |
| N | 3631 | 3631 | 3356 | 3356 | | | |
| Year FE | \checkmark | \checkmark | \checkmark | \checkmark | | | |
| Municipality FE | \checkmark | \checkmark | \checkmark | \checkmark | | | |
| Municipality controls | \checkmark | \checkmark | \checkmark | \checkmark | | | |

Table: Evaluation of the common trend

Standard errors clustered at municipality level.

Significance levels: *** 1% ** 5% * 10%.

^a We exclude transfers because not available for entrants in 2011

- Another important assumption is the absence of reverse causality
- we exclude any direct effect of expenditure on the decision to join a Municipal Union.
- To test this assumption we estimate the conditional probability to enter a Union for a municipality *i* at time *t*, given that the event has not yet occurred.
- Our hypothesis is that the coefficient of log expenditures is not significant, i.e. there is no reverse causality.
- The results of the duration model confirm our hypothesis.

Table: Estimates of the effect of the expenditures on the probability to join the union

| | Homogenous | | | | Heterogenous | | | | |
|---------------------|------------|-------------|-----------------|------------|--------------|-----------|-----------------|---------------|-------------|
| | Coeff. | s.e | Hazard Ratio | LogL | Coeff. | s.e | Hazard Ratio | LogL | P-val* |
| Log expenditures | -0.317 | (0.330) | 0.728 | -473.537 | -0.317 | (0.330) | 0.728 | -473.537 | 0.496 |
| *LR test of model | with No | rmal distri | buted het | erogeneity | against m | odel with | out contro | olling for he | terogeneity |
| All models contain | n the sam | e control | variables. | | | | | | |
| Significance levels | : *** 1% | ** 5% * | 10%. | | | | | | |

- We estimate our DiD model using a more homogeneous definition of the control group.
- We restrict the sample to the years 2001-2008
- and we exclude municipalities that never enter a Union (never treated).
- We include in CG the municipalities that join a Union between 2009 and 2011 (future treated)
 - together with all the municipalities observed in the years before entrance (within the period 2001-2008).

| Dependent variable: | Log Expenditures - 2001-2008 | | | | |
|-------------------------|------------------------------|----------------|--|--|--|
| | Full sample | Matched sample | | | |
| | (1) | (2) | | | |
| Municipal Union | -0.069*** | -0.071*** | | | |
| | (0.018) | (0.018) | | | |
| Ν | 1217 | 1201 | | | |
| Year FE | \checkmark | \checkmark | | | |
| Municipality FE | \checkmark | \checkmark | | | |
| Municipality controls | \checkmark | \checkmark | | | |
| Standard errors cluster | red at municipa | ality level. | | | |
| Significance levels: ** | | 5 | | | |
| We exclude transfers k | | | | | |
| | | | | | |
| Control group includes | s future treated | 1. | | | |

| Ferrares | Mid | rali | Ri | 770 |
|-----------|----------|-------|----|-----|
| I CITALES | VIIP | çanı, | | 120 |

- We analysed the Italian experience of Municipal Unions, using unique administrative data on the municipalities in Emilia Romagna region.
- In the period considered, 2001-2011, almost half of the municipalities form a MU.
- We employed a DiD approach combined with matching models to estimate the causal effect of being member of a Municipal Union on local expenditures.
- Participation to a U reduces total current per-capita expenditures by 5%, compared to municipalities not in U.

- Effect persistent and increases up to nine years from the participation to a U.
- Entering a U do not decrease the quality of public services.
- We conclude that MU is an efficient tool that allows municipalities to gain efficiency, in terms of a reduction in public expenditures.

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